# Section B: Assessment of California halibut from the US-Mexico border to Point Conception. 

## B1. Model

A sex-structured model with different growth, natural mortality, and selectivity for males and females developed using Stock Synthesis (Methot 2009) is used to assess California halibut from the US-Mexico border to Point Conception. Sex-specific length composition data are not available for all fisheries, are not available for the fishery-independent surveys, and the sample sizes are low for many of the fisheries for which sex data is available. Therefore, selectivity is shared among some fisheries and several assumptions are made when defining selectivity.

## B1.1. Fisheries

Seven fisheries are modeled. These fisheries were selected because they represent the majority of the fish caught and because they have length composition data. The fisheries include 1) bottom trawl; 2) single-rigged trawl; 3) trawl with a footrope diameter less than 8 inches; 4) commercial hook-and-line; 5) gill net; 6) Commercial Passenger Fishing Vessel (CPFV)/charter boats; and 7) private/rental boats. Other recreational catch (e.g., shore-based fishing) is combined with fishery 7. There is ambiguity in some of the definitions of the trawl gear and which gear is recorded on catch reports. Despite these ambiguities, the three trawl gears are kept separate to allow comparison of length composition data. There is a large amount of catch for which the fishing method is unrecorded. This catch is proportionally applied to the fisheries based on detailed investigation of the gear type for a subset of the data. The commercial catch from fisheries other than hook-and-line and trawl is added to the gill net fishery. Since, the hook-and-line and gill net gears are assumed to have the same selectivity, these other gear types are assumed to have the same selectivity as these two gears.

## B1.2. Initial conditions

The stock is modeled starting in 1971 because this is when the minimum legal size (MLS) was introduced, which would have substantially changed the behavior of the fisheries, and there is a lack of information on the amount of discarded fish, the size of discards, and the discard mortality rate. Length composition, abundance, and detailed catch data is also lacking from the early period. The historical catch data starts in 1916 when the commercial fishery was already substantial, so even if the population was modeled starting in 1916 it would have to be modeled from an exploited state. The initial age-structure of the population is constructed using a parameter to scale the average equilibrium recruitment, fishing mortality parameters for the CPFV fishery, and cohort strength deviates for ages zero to six with tapering bias correction. The initial equilibrium catch is not fit in the model due to the change in MLS. The initial condition
setup is just a way to parsimoniously parameterize the initial conditions and does not represent any particular population dynamics processes.

## B1.3. Growth and fecundity

Growth is fixed in the model using the von Bertalanffy growth curve following estimates based on recent age-length data provided by Reilly and Tanaka (Table B1.3.1). Variation of length-atage is fixed in the model as a linear function of length based on estimates of variation of length-at-age from randomly sampled fish presented in Pattison and McAllister (1990) and MacNair et al. (2001). Weight-at-length is fixed in the model based on parameters presented in Reed and MacCall (1988). Maturity-at-length is modeled using a logistic curve and fixed in the model based on data from Love and Brooks (1990). Fecundity is assumed to be proportional to weight.

## B1.4. Natural mortality

Natural mortality ( M ) is fixed at 0.2 for females and 0.3 for males based on assumptions made for summer flounder. However, recent work suggests $M$ may be higher for summer flounder than assumed in the summer flounder assessment. These values are essentially arbitrary since there are no reliable data to estimate natural mortality for California halibut. However, for this stock and other flatfish it is clear natural mortality is higher for males compared to females.

## B1.5. Recruitment

Recruitment is assumed to be independent of stock size (i.e., steepness $=1$ ). This is essentially arbitrary since there are no reliable data to estimate steepness for this stock. The sex ratio at birth is assumed to be 1:1. Recruitment deviates are estimated for all years and the standard deviation of the logarithm of the recruitment deviates is assumed equal to 1.0. This is higher than the typical 0.6 because recruitment appears to be more variable for this stock. Most of the information on annual recruitment variation comes from the length composition data. A ramped recruitment deviate bias correction factor was used with full bias correction starting in 1973 and ending in 2006.

## B1.6. Catch

Commercial catch by fishery in metric tons (Table B1.6.1; Figure B1.6.1) and recreational catch in thousands of fish (Table B1.6.1; Figure B1.6.2) are included in the model.

## B1.7. Selectivity

The Stock Synthesis double normal selectivity option is used for all fisheries, including the gill net fishery, which is assumed to have asymptotic selectivity (i.e. reaching a maximum finite level). Some of the six parameters are fixed to simplify the model assumptions due to the limited availability of composition data and the final selectivity parameter is ignored (option 999). Male selectivity is modeled as a dogleg offset of the female selectivity. Discards are
modeled using a logistic retention curve. Selectivity is length specific to better reflect the differences in growth rates between males and females.

All trawl fisheries are assumed to have the same selectivity due to low length composition sample sizes and limited years of data for the single-rigged trawl and trawl with a footrope less than 8 inches fisheries. The selectivity at the lower length bin (lower bound of 1 cm ) is assumed to be zero since trawl gear does not catch these very small fish. The length at the peak and the width of the ascending limb is fixed to represent the federal observer program discards (annual raw data was not available). Retention is fixed using a steep logistic curve with the inflection at the minimum legal size of 22 inches ( 559 mm ). All discarded fish are assumed to die. Male selectivity is estimated as an offset from females with the dogleg set at 559 mm and the offset at the dogleg and largest length bin estimated.

Hook-and-line selectivity is set equal to the gill net selectivity due to lack of length frequency data. An assumption of $10 \%$ of discarded fish dying is used based on data from mortality experiments for California halibut.

The gill net fishery selectivity is assumed to be logistic as it is difficult to estimate the model parameters unless at least one selectivity curve is asymptotic and the gill net fishery tends to catch the larger fish. The selectivity at the lower length bin is assumed to be zero. Retention is estimated using a logistic curve since there are length composition data on discards from an observer program and a management allowance for some retention of undersized fish in the commercial fishery. An assumption of $30 \%$ of discarded fish dying is used based on data from mortality experiments for California halibut. Male selectivity is estimated as an offset from females with the dogleg set at 650 mm and the offsets at the dogleg and largest length bin estimated.

The CPFV fishery selectivity at the lower length bin is assumed to be zero. The logistic retention curve is estimated since there is length composition data for both the whole catch and the retained catch. The asymptotic retention level was initially allowed to differ between 1971-1990 and 1991-2008 because of the two periods of discard data (logbooks and observer. The recreational survey estimates were considered unreliable). However, initial runs estimated retention of legal sized fish to be $100 \%$ so the asymptotic retention parameter was fixed at 1.0 for both periods. No male offset is calculated for retention since there is no sex-specific composition data and male and female small fish are assumed to behave the same. Discard mortality is assumed to be 10\% based on data for summer flounder and mortality experiments for California halibut. The male selectivity offset is assumed to be similar to that estimated for the commercial gears ( $25 \%$ of females at 650 mm and $0.03 \%$ at the largest size) because there is no sex-specific length composition data for the recreational fishery.

Due to similar length composition data and lack of length composition for the discarded fish in the other recreational fishery, the selectivity is set equal to that estimated for the CPFV fishery.

There is no sex-specific size composition data for the fishery-independent indices of abundance, so sex-specific selectivities cannot be estimated. When included in the model (during initial sensitivity runs not shown here), length based selectivities are used in an attempt to account for differences in growth between males and females, but this would not account for any agespecific differences in selectivity between males and females (e.g., due to differences in ontogenetic movement). A double normal is used to model the selectivity with selectivity at the largest size in the model set to zero. The length based selectivity assumption may be adequate for indices that relate to mostly juveniles.

The selectivity of the swept area survey, when included in the model (during initial sensitivity runs not shown here), is modeled using an age-based double normal with male offset for the peak, width of ascending limb, width of descending limb, and selectivity at final size bin.

## B1.8. Summary of data used in the assessment

The model is fit to an index of relative abundance (index) based on standardized catch-per-unit effort (CPUE) from CPFV logbooks. Only data from CPFV trips that were considered to target California halibut based on expert judgment about associated species are used to create the index, which extends from 1980 to 2010. The data was split into two inshore areas separated at Palos Verdes and an island area before standardizing and then a composite index was created using an average weighted by the number of blocks in each area. The index is fit using a lognormal likelihood function with a standard deviation for the logarithm of the index set to 0.3 (approximately equivalent to a coefficient of variation (CV)).

The model is fit to CPFV discard catch estimates from observer data (1986-1989) and logbooks (1995-2007). The discard data is fit using a normal likelihood function with standard deviations based on an assumed CV of 0.6 to limit their influence on the estimates of abundance.

The model is fit to bottom trawl and gill net length composition data for both combined sexes and for proportions of males and females that sum to one. Years with sample sizes less than 20 fish were excluded from the analysis. Data from hook-and-line, and the two other trawl methods were excluded because the samples sizes are small and it is considered better to remove fish at somewhat wrong sizes than to mix length composition samples from gears that may have different selectivities. Length composition data of retained and discarded halibut from a gill net observer program were included in the model to estimate the gill net retention curve. The model is fit to both whole catch and to retained catch length (or weight) composition data for the CPFV fishery and to retained catch length (or weight) composition for the private/rental boat fishery. Length composition data for the other recreational fisheries (shore based fisheries) is not used because the sample sizes are small and the catch by those fisheries is small. A multinomial likelihood with the recorded sample sizes is used for the likelihood function.

Other data are available and the results are compared to these data, but the data are not used to estimate the parameters in the assessment except in sensitivity analyses.

## B2. Data

## B2.1. CPFV logbook

Only data from CPFV trips that were considered to target California halibut based on expert judgment about associated species are used to create the index, which extends from 1980 to 2010. Data is also available prior to 1980, but lacks details to identify the target species. The catch of California halibut per angler hour for inshore blocks separated into north and south of Palos Verdes and blocks around islands were standardized separately for the categorical variables year, month, and block using a delta-lognormal regression (Table B2.1.1, Figures B2.1.1 and B2.1.2). The combined year effect from the binomial and lognormal components of the regression was used as the index of relative abundance. A combined index was created as the average weighted by the number of blocks in each sub-area. The index is assumed proportional to the number of fish selected by the CPFV fishery. The index was used in the assessment model. The index is fit using a lognormal likelihood function with a standard deviation for the logarithm of the index set to 0.3 (approximately equivalent to a CV).

## B2.2. California Cooperative Oceanic Fisheries Investigations

The catch of California halibut larvae per volume of water strained in the oblique tows was standardized for the categorical variables year, month, and station using a delta-lognormal regression. The data was restricted to tows conducted between Point Conception and the USMexico border and that started at 76.2 m (approximately 250 ft ) deep or less as a proxy for ocean depth and is related to where most California halibut larvae are found. Pelagic larvae occur over the shelf with greatest densities in water less than 250 ft deep and within 4 miles of shore. The combined year effect from the binomial and lognormal components of the regression was used as the index of relative abundance (Table B2.2.1, Figure B2.2.1). The index is assumed proportional to spawning biomass rather than recruitment since there is a shorter time period between spawners and eggs/larvae than there is between eggs/larvae and recruitment to the fishery. Therefore, there is a lower probability that external factors can influence the relationship between spawners and eggs/larvae. The index is highly variable from one year to the next, has high coefficients of variation, and includes several zero observations that cannot be included directly in the lognormal likelihood function commonly used to fit indices of abundance. Therefore, it is considered an unreliable index of abundance and not used in the assessment model. Similar results were obtained if the data was not restricted by depth at the start of the tow.

## B2.3. Ocean Resources Enhancement Hatchery Program gill net survey

The OREHP gill net surveys for juvenile white seabass also catch California halibut. The gill nets were set overnight and floated with the lead lines about 1 m above the bottom so they may not be an optimal halibut sampling gear. Two surveys were conducted. The California State University of Northridge (CSUN) survey (data provided by Dr. Larry G. Allen, Southern California Marine Institute) covered the northern two-thirds of the Southern California Bight from

Newport Beach to Santa Barbara including Catalina Island. The San Diego State [Hubbs-Sea World Research Institute (HSWRI)] survey (data provided by Michael A. Shane, HSWRI) covered the stations south to the US-Mexico border. HSWRI has been deploying multiple mesh gill nets since 1988, but the mesh sizes and panel lengths in the nets varied until June 1992 when the nets were standardized. Hence only data starting in June 1992 from nets with similar design (mesh and panel sizes) were used. The CSUN survey had different mesh sizes, which varied over time. Therefore, the data was used to create two indices of abundance. Each survey has associated length composition data that can be used to estimate a selectivity curve to determine the component of the population it relates to.

The catch of California halibut per length of net in the CSUN survey was standardized for the categorical variables year and month using a linear regression. The year effect was used as the index of relative abundance (Table B2.3.1, Figure B2.3.1). The data was provided with all stations combined so area or station effects could not be used in the standardization.

The catch of California halibut per length of net in the HSWRI survey was standardized for the categorical variables year, month, and site using a delta-lognormal regression. The combined year effect from the binomial and lognormal components of the regression was used as the index of relative abundance (Table B2.3.1, Figure B2.3.1). Jackknife estimates of error were calculated for the HSWRI survey.

Each of these indices is only from a part of the coast, there is no sex-specific length composition information, they were not designed to capture halibut (lead lines about 1 m above the bottom), the CSUN survey has inconsistent mesh sizes, and the CV for the HSWRI survey was large in many years. Therefore, they are considered unreliable indices of abundance and not used in the assessment model. Length composition data is available for the surveys (Table B2.8.3).

## B2.4. Impingement and associated trawl survey

Monitoring has been carried out since 1972 at all electricity generating coastal facilities built by Southern California Edison. Three operational survey types are used: normal operation, heat treatment, and fish chase. A trawl survey is also conducted. The impingement number per unit of volume data (provided by Eric Miller), which includes both heat treatment and fish chase surveys, was standardized for the categorical variables year, month, survey type, and station using a delta-lognormal regression. The trawl data was standardized for the categorical variables year, month, site, and station using a delta-lognormal regression. The combined year effect from the binomial and lognormal components of the regression was used as the index of relative abundance (Table B2.4.1, Figure B2.4.1). A jackknife procedure was used to estimate standard errors, but it did not work for the trawl data. Each survey has associated length composition data that can be used to estimate a selectivity curve to determine the component of the population it relates to. The indices represent a small spatial area, there is no sex-specific length composition information, and the CV for the impingement index is moderately high for
many years. Therefore, they are considered unreliable indices of abundance and not used in the assessment model.

## B2.5. Sanitation district data

Records of California halibut catch and length composition data were obtained from three sanitation districts (Table B2.5.1, Figure B2.5.1). None of these had details about the trawl tows that did not catch California halibut. Therefore, indices of abundance were calculated by assuming the stations were sampled the same number of times each year. The sanitation data is from a limited area, there is no sex-specific length composition information, and therefore they are not used in the assessment.

## B2.6. Trawl logbook data

Attempts were made to standardize the catch of California halibut per tow from the trawl logbook data using the categorical variables year, month, block, and depth using a deltalognormal regression. The data were aggregated by these categories for analysis. The data was limited to trips that caught California halibut and/or species associated with California halibut (based on expert opinion). The analysis was unstable with wide fluctuations in the index of abundance depending on the covariates included. The regression did not converge with some combinations of covariates. Therefore, the results did not appear to be reasonable. Therefore, unstandardized data (using tows targeting halibut) was considered as an index of abundance and assumed proportional to the fish vulnerable to the bottom trawl fishery (Table B2.6.1., Figure B2.6.1). Due to the issues in standardizing the index it is considered an unreliable index of abundance and not used in the assessment model.

## B2.7. Swept area trawl estimate of abundance

A trawl survey was carried out in the southern area during 1994 (Domeier). The data was used to generate a swept area estimate of abundance. The fish were weighed after their gonads were removed so the estimates of number of fish are more reliable than estimates of biomass. The estimated number of fish in the southern area, including islands, was $3,862,104$ with $90 \%$ confidence intervals $(\mathrm{Cl}) \pm 712,740$. Sex-specific age-length data is also available for this survey. However, the codend mesh size of the 400-mesh Eastern trawl used was 85 mm ( 3.34 inches) and was likely responsible for the scarcity of larger, older fish and the relatively low proportion of aged halibut that were females (31 per cent). Swept area survey estimates are notoriously biased due to invalid assumptions about the proportion of fish within the trawls path that are caught and the distribution of fish outside the trawled areas. It is not possible to estimate the catchability coefficient from a single biomass estimate. Therefore, this abundance estimate was not used in the model.

## B2.8. Length composition data

Length composition data is available from several sources. Commercial length composition by gear and sex come from California Department of Fish and Game's (CDFG) State Finfish Management Project (2007-2010) and market sampling (1983-2006) databases (Table B2.8.1, Figure B2.8.1). The length composition data on the retained and discarded halibut that are available from the gill net observer program (1983-1989) were included in the model to estimate the gill net retention curve. Some data is not available by sex. Recreational length composition data (unsexed) of the whole catch (retained plus discarded) in the CPFV fishery comes from the observer sampling (1975-1978 and 1986-1989) and retained catch comes from the Recreational Fisheries Information Network (RecFIN) (1993-2010) database (Table B2.8.2, Figure B2.8.2). Weight composition data from the retained catch for the recreational fisheries is also available from the RecFIN (1980-1989) database. Length composition data is also available for the fishery-independent surveys (impingement, impingement trawls, sanitation trawls, and gill net surveys; Table B2.8.3, Figure B2.8.3). The composition data were reweighted by estimating the effective sample size using a Michaelis-Menton (Beverton-Holt) equation in a single iteration (Figure B2.8.4).

## B2.9. Age composition data

Randomly collected age data is available from 1985-1988 (Sunada et al. 1990; Table B2.9.1). It is from both gill net and trawl fisheries combined. Therefore, it is not used in the assessment.

Sex-specific age composition data is available for the trawl survey (MacNair et al. 2001; Table B2.9.2). Since the trawl survey is not used in the assessment model and the length-at-age data may be biased due to selectivity of the trawl codend mesh size, the corresponding age data is not used.

## B2.10. Discards

Estimates of recreational discard rates are available from RecFIN, CPFV observers, and CPFV logbooks (Table B2.10.1, Figure B2.10.1). The total discards can be calculated by applying these ratios to the retained catch (Table B2.10.2). Initial analyses estimated the retention rate of legal sized fish to be $100 \%$, so the asymptotic retention value was set at 1.0.

Commercial discards are available from the federal observer program on halibut trawl trips, both limited entry and open access (Table B2.10.3, Figure B2.10.2). However, most of the data is from north of Point Conception and samples sizes are low. Limited data is available on the size composition of the discards (Figure B2.10.3.). An annual series of discard rates is available (e.g., see Bellman et al. 2010). Discard length composition data is available for the gill net fishery from the gill net observer program (see above).

## B3. Results

## B3.1. Model fit

Visually, the stock assessment model provides a poor fit to the CPFV index of relative abundance due to temporal correlation in the residual pattern (Figure B3.1.1a). However, the fit is quite good statistically and the root mean square error (RMSE) (0.22) of the fit to the CPFV index is lower than the assumed standard error (SE) for this index (0.3), which can be seen visually as the predicted time series lies within the $95 \%$ confidence intervals of the observed index points for all years. The low RMSE is due to the lack of variation in the observed index.

The model fits the length composition data (Appendix B.1), and the CPFV discard data (Figure B3.1.2) reasonably well. The reweighted sample sizes for the length composition data are all smaller than the input sample size (Figure B2.8.4). The model provides poor fits to the trawl logbook and CalCOFI indices of abundance (Figures B3.1.1b, c; note they are not used in estimating the model parameters). However, the trend in the trawl logbook index is similar to that estimated by the model. The fit to the other indices of abundance are not meaningful since the selectivities are not estimated in the base case assessment.

## B3.2. Estimates

The stock is estimated to be depleted with the start of year 2011 spawning biomass estimated to be only $14 \%$ of the unexploited spawning biomass (Figure B3.3.1). This is not completely unexpected since the population was already exploited at the start of the modeling period in 1971 ( $16 \%$ of the unexploited level) and the CPFV index of relative abundance is relatively flat. There was substantial catch prior to 1971, but changes in the management (e.g., minimum legal size) and lack of data prior to 1971 prevent adequate assessment of the abundance prior to 1971. The assessment is able to estimate the depletion level in 1971 by calculating the abundance at age that is consistent with the length composition data seen in the fisheries compared to what would be expected in the absence of fishing. It is not unusual for a fish stock to be sustainably fished at high fishing levels and low abundance levels for decades. There has been a series of low recruitments in recent years (since 1999; Figure B3.2.2). Recruitment does not appear to be related to spawning stock size (Figure B3.2.3).

The estimated selectivity curves appear reasonable (Figure B3.2.4). Male selectivity is estimated to be substantially less than female selectivity.

## B3.3. Management quantities

Maximum sustainable yield (MSY) is estimated to occur at a very low fraction of the unexploited spawning biomass (Table B3.3.1). Therefore, even though the population is estimated to be depleted, it is still estimated to be above the spawning biomass level that would produce MSY and the fishing mortality is lower than the level that would produce MSY. This is partly due to the assumption recruitment is independent of stock size, at least at the abundance levels observed, which is not contradicted by the data. In other words, halibut as a species, like many other flatfishes, are prolific enough, and have a high reproductive potential, such that when environmental conditions are favorable, biomass can increase relatively quickly in a short time frame. MSY and related quantities are also dependent on the size of fish caught, natural
mortality, and growth. There is substantial uncertainty about many of the biological and fishing processes including the stock-recruitment relationship, natural mortality, growth, and the survival of discarded fish. Despite the resilience of flatfish and the fact that California halibut have sustained high exploitation rates for several decades, the uncertainty in the biological and fishing processes and the recent series of low recruitments indicate that management action may be needed to reduce the risk of fishery collapse.

## B3.4. Sensitivity analyses

Several analyses were carried out to evaluate the sensitivity of the results to model assumptions: 1) using the longer time frame aggregated CPFV CPUE index of abundance, 2) steepness of the stock-recruitment relationship equal to $0.75,3$ ) the average length of age 20 females equal to $100 \mathrm{~cm}, 4$ ) natural mortality for females and males fixed at 0.3 and 0.4, respectively, and 5) not reweighting the composition data (Table B3.4.1). The RMSE for the aggregated CPFV CPUE (0.39) is much worse than for the detailed logbook targeting based index. However, visually the model appears to fit the index better and shows a decline in recent years consistent with the estimated low recruitments (Figure B3.4.1). The likelihood is worse when the steepness of the stock-recruitment relationship is fixed at 0.75 and the population is more depleted. Both reduced average length at age 20 and increased natural mortality have better likelihood values and lower depletion levels. Using the original samples sizes for the composition data has a higher RMSE for the CPFV index of abundance and the population is slightly less depleted.

## B3.5. Discussion

The stock assessment estimates the California halibut population south of Point Conception has been depleted during the whole time frame modeled (1971-2011). The depletion level is somewhat sensitive to uncertainties in several important biological processes (e.g., growth and natural mortality) and data (e.g., method used to standardize the CPFV index of abundance). CPFV CPUE data suggests the status of the stock may differ among locations in the southern region and it is unclear if these should be treated as separate stocks.

The stock assessment results are somewhat consistent with the alternative data sets. Although the CaICOFI data was not used in the model, the CalCOFI index was low in the late 1990s and early 2000s suggesting low egg production or some oceanographic process may have generated low recruitment (Figure B3.1.1c). The gill net surveys, which catch small halibut, were low in the north in the late 1990s and early 2000s, but not in the south (Figure B2.3.1). The raw bottom trawl catch per tow is consistent with the biomass trend (Figure B3.1.1b). The impingement heat treatment and fish chase (HTFC) and trawl surveys, which catch small halibut, were also generally lower in the early 2000s (Figure B2.4.1). The sanitation surveys, which also catch small fish, were lower in the mid 2000s (Figure B2.5.1). These all generally support a low recruitment level in the late 1990s and early 2000s.

## Tables and Figures

Table B1.3.1. Growth, variation in length-at-age, length-weight, and maturity parameters used in the assessment. K, L1 and L20 are the parameters for the version of the von Bertalanffy growth equation used in stock synthesis. Lsd1 and Lsd20 are the parameters of the linear relationship between the standard deviation of the variation in length at age and age. The quantities $a$ and $b$ are the parameters of the length-weight relationship. The quantities alpha and beta are the parameters of the logistic maturity curve and L50 is the length at 50\% maturity.

| Growth | Females | Males |
| :--- | ---: | ---: |
| K | 0.095 | 0.112 |
| L1 | 22.1 | 15.3 |
| L20 | 114.5 | 107.3 |
| Variation |  |  |
| Lsd1 | 17.6 | 13.8 |
| Lsd20 | 161.9 | 140.3 |

Length weight

| a | $7.77 \mathrm{E}-06$ | $9.22 \mathrm{E}-06$ |
| :--- | ---: | ---: |
| a | 3.0496 | 3.0165 |
| Maturity |  |  |
| alpha | -0.15 | -0.34 |
| beta | 7.02 | 7.77 |
| L50 | 47.1 | 22.7 |

Table B1.6.1. Catch for each of the fisheries used in the assessment of the stock south of Point Conception. Catch is in kg for commercial fisheries (Comm) and in number of fish for the recreational fisheries (Rec).

| Year | Comm bottom trawl | Comm <br> single <br> rigged <br> trawl | Comm <br> trawl <br> with <br> footrope <br> diameter <br> < 8 <br> inches | Comm hook-and-line | Comm gill net | Rec CPFV | Other recreational |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1971 | 30 | 0 | 0 | 1,251 | 55,745 | 22 | 82,269 |
| 1972 | 15,456 | 0 | 0 | 2,488 | 40,775 | 13,915 | 54,913 |
| 1973 | 24,075 | 660 | 0 | 3,972 | 44,000 | 16,162 | 66,304 |
| 1974 | 24,899 | 0 | 0 | 2,729 | 33,245 | 16,992 | 72,530 |
| 1975 | 18,417 | 0 | 0 | 3,539 | 66,383 | 16,978 | 75,485 |
| 1976 | 3,274 | 0 | 0 | 1,983 | 89,245 | 18,394 | 85,269 |
| 1977 | 0 | 0 | 0 | 4,106 | 77,730 | 13,868 | 67,112 |
| 1978 | 1,145 | 8 | 0 | 4,758 | 107,130 | 10,955 | 55,411 |
| 1979 | 1,538 | 0 | 0 | 4,300 | 152,779 | 10,562 | 55,909 |
| 1980 | 1,117 | 0 | 0 | 2,436 | 57,250 | 8,007 | 107,440 |
| 1981 | 501 | 0 | 0 | 4,994 | 259,933 | 16,495 | 54,098 |
| 1982 | 1,450 | 0 | 0 | 2,634 | 229,117 | 15,280 | 68,831 |
| 1983 | 2,491 | 0 | 0 | 1,902 | 148,949 | 3,662 | 26,822 |
| 1984 | 4,591 | 0 | 0 | 2,605 | 193,923 | 2,011 | 32,524 |
| 1985 | 14,963 | 0 | 0 | 885 | 307,947 | 11,539 | 57,164 |
| 1986 | 22,567 | 0 | 0 | 3,577 | 223,533 | 11,014 | 109,374 |
| 1987 | 27,758 | 27 | 0 | 6,729 | 230,296 | 26,825 | 135,664 |
| 1988 | 16,182 | 5 | 0 | 9,507 | 220,219 | 18,049 | 67,562 |
| 1989 | 12,198 | 0 | 0 | 4,950 | 196,432 | 22,420 | 91,224 |
| 1990 | 20,063 | 0 | 0 | 5,040 | 187,683 | 12,152 | 67,424 |
| 1991 | 18,300 | 0 | 0 | 6,932 | 219,303 | 11,004 | 61,054 |
| 1992 | 23,383 | 164 | 0 | 3,138 | 126,953 | 5,041 | 27,968 |
| 1993 | 11,668 | 543 | 0 | 3,596 | 110,854 | 7,432 | 27,643 |
| 1994 | 10,751 | 19,686 | 0 | 3,805 | 50,085 | 13,699 | 39,712 |
| 1995 | 6,502 | 44,912 | 0 | 8,399 | 71,718 | 8,897 | 61,328 |
| 1996 | 2,782 | 59,351 | 0 | 6,686 | 86,625 | 10,468 | 49,664 |
| 1997 | 2,384 | 34,556 | 0 | 10,406 | 127,192 | 4,853 | 40,753 |
| 1998 | 24,949 | 35,870 | 0 | 8,739 | 105,065 | 4,657 | 57,689 |
| 1999 | 26,689 | 42,311 | 0 | 15,890 | 173,664 | 16,656 | 88,556 |
| 2000 | 6,517 | 25,891 | 0 | 16,676 | 156,946 | 10,955 | 96,319 |
| 2001 | 46,298 | 24,876 | 0 | 14,257 | 138,008 | 8,616 | 82,485 |
| 2002 | 61,096 | 23,642 | 414 | 20,404 | 111,356 | 3,379 | 113,070 |
| 2003 | 24,314 | 27,310 | 18 | 14,402 | 82,620 | 6,074 | 85,875 |
| 2004 | 1,552 | 47,861 | 20 | 21,407 | 83,576 | 3,325 | 25,284 |


| 2005 | 320 | 34,848 | 786 | 16,591 | 48,583 | 4,218 | 22,959 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 2006 | 367 | 32,354 | 17 | 19,782 | 47,824 | 912 | 31,718 |
| 2007 | 1,937 | 12,452 | 21,925 | 18,542 | 44,827 | 918 | 19,159 |
| 2008 | 541 | 1,186 | 37,486 | 19,394 | 50,959 | 784 | 19,154 |
| 2009 | 677 | 5,251 | 41,268 | 19,835 | 54,472 | 1,145 | 16,252 |
| 2010 | 2,093 | 5,226 | 41,859 | 9,743 | 45,026 | 775 | 17,331 |

Table B2.1.1. Indices of relative abundance for the CPFV logbook data. Data from CPFV trips that were considered to target California halibut were standardized for the categorical variables year, month, and block using a delta-lognormal regression. Obs is the number of trips used to generate the index in that year. The combined index was created as the average weighted by the number of blocks in each sub-area. The indices are assumed proportional to the number of fish selected by the CPFV fishery and the combined index is used in the stock assessment model.

|  | Islands (30 blocks) |  | North (19 blocks) |  | South (16 blocks) |  | Combined |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Index | Obs | Index | Obs | Index | Obs | Index |
| 1980 | 0.017115 | 305 | 0.035399 | 872 | 0.018452 | 483 | 0.022789 |
| 1981 | 0.015183 | 337 | 0.030275 | 1182 | 0.017866 | 511 | 0.020255 |
| 1982 | 0.016472 | 388 | 0.033123 | 1222 | 0.020209 | 800 | 0.022259 |
| 1983 | 0.012964 | 213 | 0.024025 | 367 | 0.016575 | 438 | 0.017086 |
| 1984 | 0.011143 | 199 | 0.023579 | 519 | 0.015963 | 271 | 0.015965 |
| 1985 | 0.01294 | 324 | 0.03183 | 915 | 0.0205 | 547 | 0.020323 |
| 1986 | 0.01695 | 292 | 0.031045 | 981 | 0.018236 | 473 | 0.021387 |
| 1987 | 0.016431 | 191 | 0.031139 | 1010 | 0.017472 | 489 | 0.020987 |
| 1988 | 0.016061 | 368 | 0.037645 | 1233 | 0.016629 | 858 | 0.02251 |
| 1989 | 0.017003 | 287 | 0.037751 | 967 | 0.019839 | 746 | 0.023766 |
| 1990 | 0.012096 | 332 | 0.031195 | 940 | 0.019556 | 767 | 0.019515 |
| 1991 | 0.014484 | 357 | 0.030539 | 929 | 0.019893 | 693 | 0.020508 |
| 1992 | 0.012773 | 193 | 0.023876 | 590 | 0.020425 | 322 | 0.017902 |
| 1993 | 0.012671 | 252 | 0.034648 | 651 | 0.020883 | 477 | 0.021117 |
| 1994 | 0.013861 | 292 | 0.033804 | 506 | 0.018499 | 677 | 0.020832 |
| 1995 | 0.013734 | 478 | 0.035383 | 848 | 0.016999 | 1295 | 0.020866 |
| 1996 | 0.017066 | 460 | 0.032702 | 1052 | 0.015085 | 1523 | 0.021149 |
| 1997 | 0.018835 | 609 | 0.031563 | 1572 | 0.013466 | 1593 | 0.021234 |
| 1998 | 0.016078 | 527 | 0.03525 | 1064 | 0.015656 | 1138 | 0.021579 |
| 1999 | 0.018712 | 1006 | 0.042284 | 1125 | 0.024838 | 1393 | 0.02711 |
| 2000 | 0.019247 | 1070 | 0.037333 | 1280 | 0.017214 | 1459 | 0.024033 |
| 2001 | 0.023191 | 1011 | 0.034015 | 1063 | 0.012601 | 1531 | 0.023748 |
| 2002 | 0.020954 | 746 | 0.03518 | 639 | 0.012594 | 1201 | 0.023054 |
| 2003 | 0.019786 | 676 | 0.035241 | 581 | 0.014642 | 1051 | 0.023038 |
| 2004 | 0.01973 | 302 | 0.034929 | 365 | 0.01362 | 731 | 0.022669 |
| 2005 | 0.018311 | 432 | 0.029208 | 287 | 0.014332 | 719 | 0.020517 |
| 2006 | 0.020774 | 610 | 0.035425 | 242 | 0.014879 | 799 | 0.023605 |
| 2007 | 0.022194 | 480 | 0.030721 | 117 | 0.013107 | 736 | 0.02245 |
| 2008 | 0.018328 | 573 | 0.027432 | 155 | 0.015126 | 775 | 0.020201 |
| 2009 | 0.020833 | 552 | 0.027779 | 89 | 0.014904 | 519 | 0.021404 |
| 2010 | 0.018569 | 316 | 0.030412 | 95 | 0.018278 | 384 | 0.021959 |

Table B2.2.1. Indices of relative abundance for the CaICOFI data. CV is the coefficient of variation.

| Year | Index | CV | Tows | Year | Index | CV | Tows |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1972 | 0 | NA | 12 | 1992 | 0.038693 | 0.80 | 19 |
| 1974 | 0.018859 | 0.62 | 27 | 1993 | 0.03897 | 0.56 | 22 |
| 1975 | 0.012978 | 0.35 | 146 | 1994 | 0.029065 | 0.75 | 21 |
| 1977 | 0.466019 | 0.86 | 8 | 1995 | 0.023834 | 0.78 | 20 |
| 1978 | 0.040799 | 0.39 | 53 | 1996 | 0.04538 | 0.54 | 21 |
| 1979 | 0 | NA | 9 | 1997 | 0.014875 | 0.89 | 22 |
| 1980 | 0 | NA | 11 | 1998 | 0.030408 | 0.45 | 32 |
| 1981 | 0.076255 | 0.62 | 25 | 1999 | 0 | NA | 22 |
| 1982 | 0 | NA | 2 | 2000 | 0.006177 | 0.80 | 15 |
| 1983 | 0 | NA | 3 | 2001 | NA | NA | 19 |
| 1984 | 0.009529 | 0.86 | 34 | 2002 | 0 | NA | 20 |
| 1985 | 0.067349 | 0.87 | 22 | 2003 | 0 | NA | 18 |
| 1986 | 0.070915 | 0.72 | 25 | 2004 | 0.030546 | 0.51 | 36 |
| 1987 | 0.009168 | 0.77 | 22 | 2005 | 0.008459 | 0.52 | 52 |
| 1988 | 0.006997 | 0.97 | 21 | 2006 | 0.012075 | 0.48 | 49 |
| 1989 | 0.03171 | 0.52 | 22 | 2007 | NA | NA | 42 |
| 1990 | 0.018648 | 0.72 | 23 | 2008 | 0 | NA | 12 |
| 1991 | 0.011487 | 0.68 | 21 |  |  |  |  |

Table B2.3.1. Indices of relative abundance for the gill net survey data. CV is the coefficient of variation.

|  | CSUN |  | HSWRI |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Index | Records | Index | Records | CV |
| 1992 |  |  | 1.682454 | 29 | 1.22 |
| 1993 |  |  | 2.699604 | 48 | 0.80 |
| 1994 |  |  | 3.861616 | 45 | 0.41 |
| 1995 | 1 | 5 |  |  |  |
| 1996 | 1.178555 | 5 | 3.836322 | 18 | 0.31 |
| 1997 | 0.78079 | 4 | 2.802623 | 36 | 0.27 |
| 1998 | 0.593652 | 4 | 1.607183 | 36 | 0.76 |
| 1999 | 0.384474 | 4 | 2.015596 | 34 | 0.54 |
| 2000 | 0.332923 | 4 | 2.988562 | 36 | 0.32 |
| 2001 | 0.525439 | 4 | 2.849381 | 37 | 0.51 |
| 2002 | 0.34695 | 4 | 2.339514 | 36 | 0.57 |
| 2003 | 0.19787 | 4 | 3.06859 | 36 | 0.58 |
| 2004 | 0.154064 | 4 | 2.495822 | 35 | 0.37 |
| 2005 | 0.824833 | 2 | 2.522458 | 30 | 0.80 |
| 2006 | 0.891373 | 2 | 2.323233 | 11 | 0.35 |
| 2007 | 0.463029 | 3 | 2.408852 | 27 | 0.39 |
| 2008 | 0.446826 | 2 | 2.446496 | 15 | 0.28 |

Table B2.4.1. Indices of relative abundance for the impingement and related trawl survey data. "Records" refers to the number of samples or tows. CV is the coefficient of variation.

| Year | Impingement |  |  | Trawl |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Index | Records | CV | Index | Records |
| 1971 |  |  |  | 0.00029 | 13 |
| 1972 | 0.004793 | 38 | 0.42 | 0.000952 | 35 |
| 1973 | 0.004449 | 35 | 0.45 |  |  |
| 1974 | 0.010468 | 44 | 0.36 |  |  |
| 1975 | 0.010419 | 42 | 0.36 | 0.003932 | 12 |
| 1976 | 0.002647 | 62 | 0.52 | 0.005684 | 6 |
| 1977 | 0.006571 | 43 | 0.43 | 0.211718 | 6 |
| 1978 | 0.014879 | 39 | 0.42 | 0.060024 | 88 |
| 1979 | 0.022858 | 60 | 0.39 | 0.067532 | 74 |
| 1980 | 0.037106 | 51 | 0.27 | 0.194794 | 174 |
| 1981 | 0.021018 | 53 | 0.33 | 0.213707 | 122 |
| 1982 | 0.032267 | 48 | 0.26 | 0.060432 | 675 |
| 1983 | 0.031276 | 48 | 0.37 | 0.046349 | 651 |
| 1984 | 0.012374 | 41 | 0.43 | 0.050408 | 619 |
| 1985 | 0.00948 | 56 | 0.30 | 0.064645 | 672 |
| 1986 | 0.024196 | 56 | 0.37 | 0.07943 | 708 |
| 1987 | 0.007125 | 47 | 0.45 | 0.066321 | 48 |
| 1988 | 0.007863 | 45 | 0.39 | 0.064792 | 142 |
| 1989 | 0.004012 | 51 | 0.42 | 0.036098 | 6 |
| 1990 | 0.005151 | 57 | 0.36 | 0.06511 | 46 |
| 1991 | 0.002992 | 45 | 0.49 | 0.16979 | 33 |
| 1992 | 0.005175 | 72 | 0.40 | 0.034103 | 46 |
| 1993 | 0.008333 | 53 | 0.32 | 0.035666 | 46 |
| 1994 | 0.007652 | 66 | 0.30 | 0.032544 | 46 |
| 1995 | 0.016498 | 50 | 0.32 | 0.066913 | 74 |
| 1996 | 0.00909 | 63 | 0.32 | 0.062444 | 78 |
| 1997 | 0.014222 | 57 | 0.28 | 0.047757 | 97 |
| 1998 | 0.028455 | 54 | 0.26 | 0.058338 | 78 |
| 1999 | 0.021596 | 35 | 0.44 | 0.054064 | 117 |
| 2000 | 0.002396 | 56 | 0.40 | 0.028838 | 112 |
| 2001 | 0.006602 | 60 | 0.31 | 0.040556 | 118 |
| 2002 | 0.007913 | 55 | 0.34 | 0.068121 | 112 |
| 2003 | 0.001797 | 49 | 0.43 | 0.03551 | 85 |
| 2004 | 0.005297 | 40 | 0.50 | 0.032246 | 117 |
| 2005 | 0.017568 | 49 | 0.34 | 0.067046 | 112 |
| 2006 | 0.005159 | 33 | 0.51 | 0.084118 | 112 |
| 2007 | 0.004783 | 50 | 0.45 | 0.062724 | 118 |
| 2008 | 0.00387 | 43 | 0.45 | 0.03533 | 106 |


| 2009 | 0.001232 | 35 | 0.62 | 0.020396 | 118 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 2010 |  |  |  | 0.013234 | 38 |

Table B2.5.1. Indices of relative abundance for the sanitation district trawl data. LACSD is the Los Angeles County Sanitation District, OCSD is the Orange County Sanitation District.

| Year | LACSD | OCSD | LA2 |
| ---: | ---: | ---: | :--- |
| 1972 | 3 |  |  |
| 1974 | 6 |  |  |
| 1975 | 1 |  |  |
| 1976 | 4 |  |  |
| 1977 | 23 |  |  |
| 1978 | 5 |  |  |
| 1979 | 31 |  |  |
| 1980 | 16 |  |  |
| 1981 | 66 |  |  |
| 1982 | 71 |  |  |
| 1983 | 90 |  |  |
| 1984 | 54 |  |  |
| 1985 | 113 | 4 |  |
| 1986 | 62 | 2 |  |
| 1987 | 76 | 5 |  |
| 1988 | 85 | 14 |  |
| 1989 | 59 | 11 | 24 |
| 1990 | 65 | 14 | 57 |
| 1991 | 41 | 7 | 86 |
| 1992 | 54 | 13 | 8 |
| 1993 | 64 | 4 | 48 |
| 1994 | 45 | 13 | 81 |
| 1995 | 76 | 1 | 109 |
| 1996 | 85 | 1 | 75 |
| 1997 | 105 | 9 | 70 |
| 1998 | 124 | 12 | 61 |
| 1999 | 127 | 7 | 64 |
| 2000 | 143 | 13 | 55 |
| 2001 | 105 | 21 | 56 |
| 2002 | 46 | 6 | 65 |
| 2003 | 4 | 1 | 52 |
| 2004 | 30 | 1 | 58 |
| 2005 | 28 | 1 | 31 |
| 2006 | 32 | 7 | 29 |
| 2007 | 23 | 3 | 48 |
| 2008 | 31 | 4 | 31 |
| 2009 |  | 5 | 11 |
|  |  |  |  |

Table B2.6.1. Indices of relative abundance for the trawl logbook data.

| Year | Index | Records | Year | Index | Records | Year | Index | Records |
| ---: | :--- | ---: | ---: | :--- | ---: | ---: | ---: | ---: |
| 1981 | 13.13068 | 56 | 1991 | 2.173588 | 28 | 2000 | 8.213026 | 226 |
| 1982 | 1.287888 | 92 | 1992 | 2.003135 | 30 | 2001 | 12.75801 | 307 |
| 1983 | 0.117583 | 30 | 1993 | 1.261276 | 62 | 2002 | 9.819932 | 147 |
| 1984 | 0 | 4 | 1994 | 9.765941 | 82 | 2003 | 9.587504 | 243 |
| 1985 | 3.755854 | 89 | 1995 | 8.999438 | 151 | 2004 | 12.34795 | 211 |
| 1986 | 3.986667 | 25 | 1996 | 20.457 | 110 | 2005 | 5.514078 | 195 |
| 1987 | 9.615522 | 35 | 1997 | 12.16804 | 317 | 2006 | 6.474595 | 152 |
| 1988 | 2.59459 | 45 | 1998 | 18.1296 | 322 | 2007 | 1.834771 | 136 |
| 1989 | 0.684957 | 102 | 1999 | 15.68217 | 260 | 2008 | 6.954767 | 135 |
| 1990 | 1.051019 | 56 |  |  |  |  |  |  |

Table B2.8.1. Sample size (number of fish measured) for commercial length composition data.

|  | Bottom trawl |  | Hook-and-line |  | Gill net |  |  |  | Singlerigged trawl |  | Trawl with footrope diameter < 8 inches |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Retained <br> Sex <br> aggregated | Sex specific | Retained <br> Sex <br> aggregated | Sex <br> Specific | catch <br> Sex <br> aggregated | Sex specific | Retained <br> Sex <br> aggregated | Sex specific | Retained <br> Sex <br> aggregated | Sex specific | Retained <br> Sex <br> aggregated | Sex specific |
| 1975 |  |  |  |  |  |  |  |  |  |  |  |  |
| 1976 |  |  |  |  |  |  |  |  |  |  |  |  |
| 1977 |  |  |  |  |  |  |  |  |  |  |  |  |
| 1978 |  |  |  |  |  |  |  |  |  |  |  |  |
| 1980 |  |  |  |  |  |  |  |  |  |  |  |  |
| 1981 |  |  |  |  |  |  |  |  |  |  |  |  |
| 1982 |  |  |  |  |  |  |  |  |  |  |  |  |
| 1983 | 215 |  |  |  | 652 |  | 594 |  |  |  |  |  |
| 1984 | 538 |  | 25 |  | 1542 |  | 2000 |  |  |  |  |  |
| 1985 | 807 | 6 |  |  | 1525 | 245 | 2705 | 782 |  |  |  |  |
| 1986 | 69 | 166 |  | 24 | 1284 | 666 | 2097 | 4835 |  |  |  |  |
| 1987 | 50 | 497 | 33 | 47 | 760 | 868 | 1988 | 3604 |  |  |  |  |
| 1988 | 17 | 391 | 32 | 175 | 802 | 164 | 1618 | 4099 |  |  |  |  |
| 1989 | 17 | 339 | 5 | 81 | 191 | 25 | 992 | 4349 |  |  |  |  |
| 1990 | 2 | 92 |  | 7 |  |  | 354 | 1508 |  |  |  |  |
| 1991 | 3 | 33 | 14 | 1 |  |  | 417 | 943 |  |  |  |  |
| 1992 | 30 | 95 | 14 |  |  |  | 1748 | 1694 |  |  |  |  |
| 1993 | 8 |  |  |  |  |  | 1496 | 437 |  |  |  |  |
| 1994 | 15 |  | 4 | 3 |  |  | 73 | 3 | 26 | 18 |  |  |
| 1995 | 227 |  |  |  |  |  | 49 | 6 | 15 |  |  |  |
| 1996 | 64 |  | 2 |  |  |  | 136 | 4 | 56 |  |  |  |
| 1997 |  |  |  |  |  |  | 1 | 1 |  |  |  |  |


| 1999 |  |  |  |  | 2 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2000 |  |  | 1 |  | 24 |  |  |  |  |  |
| 2001 |  |  | 39 |  | 15 |  | 5 |  |  |  |
| 2002 |  |  | 43 |  | 491 | 1 | 15 |  |  |  |
| 2003 |  |  | 38 |  | 74 | 5 | 46 |  |  |  |
| 2004 |  |  |  |  | 38 |  | 1 |  |  |  |
| 2005 |  |  | 7 |  | 23 |  |  |  |  |  |
| 2006 |  |  | 7 |  |  |  |  |  |  |  |
| 2007 | 4 | 8 | 8 |  | 2 | 1 | 6 |  | 58 | 1 |
| 2008 |  |  | 132 | 3 | 110 | 32 | 17 | 20 | 300 | 96 |
| 2009 | 32 |  | 54 | 2 | 27 | 94 | 12 | 20 | 67 | 108 |
| 2010 | 29 |  | 24 |  | 43 | 50 | 36 |  | 4 | 1 |

Table B2.8.2. Sample size (number of fish measured or weighed) for recreational length and weight composition data. Males and females are aggregated.


Table B2.8.3. Sample size (number of fish measured) for survey length composition data.

|  | Gill net Survey HSWRI | Gill net Survey CSUN | Impingement HTFC | Impingement Trawl |
| :---: | :---: | :---: | :---: | :---: |
| Year |  |  |  |  |
| 1971 |  |  |  |  |
| 1972 |  |  |  |  |
| 1973 |  |  |  |  |
| 1974 |  |  | 1 |  |
| 1975 |  |  | 4 |  |
| 1976 |  |  |  |  |
| 1977 |  |  |  |  |
| 1978 |  |  | 1 |  |
| 1979 |  |  |  |  |
| 1980 |  |  |  |  |
| 1981 |  |  | 197 |  |
| 1982 |  |  | 79 |  |
| 1983 |  |  | 90 |  |
| 1984 |  |  | 72 |  |
| 1985 |  |  | 46 |  |
| 1986 |  |  | 80 |  |
| 1987 |  |  | 26 |  |
| 1988 |  |  | 27 |  |
| 1989 |  |  | 24 |  |
| 1990 |  |  | 22 |  |
| 1991 |  |  | 10 |  |
| 1992 | 43 |  | 20 |  |
| 1993 | 120 |  | 30 |  |
| 1994 | 157 |  | 75 |  |
| 1995 |  | 119 | 102 |  |
| 1996 | 58 | 139 | 63 |  |
| 1997 | 85 | 115 | 73 |  |
| 1998 | 40 | 107 | 88 |  |
| 1999 | 53 | 69 | 40 |  |
| 2000 | 95 | 66 | 11 |  |
| 2001 | 69 | 92 | 21 |  |
| 2002 | 67 | 70 | 33 |  |
| 2003 | 70 | 35 | 9 |  |
| 2004 | 73 | 33 | 23 | 67 |
| 2005 | 56 | 74 | 43 | 134 |
| 2006 | 18 | 77 | 26 | 197 |
| 2007 | 60 | 63 | 30 | 159 |


| 2008 | 32 | 29 | 22 | 88 |
| :--- | :--- | :--- | :--- | :--- |
| 2009 |  |  | 12 | 62 |
| 2010 |  |  |  | 17 |

Table B2.9.1. Percent age composition data from Sunada et al. (1990) $\mathrm{N}=$ number of fish aged.

| N |  | 95 | 344 | 565 | 330 |
| :--- | ---: | ---: | ---: | ---: | ---: |
|  |  |  |  |  |  |
| Age |  | 1985 | 1986 | 1987 | 1988 |
|  | 4 | 8.4 | 4.4 | 0.3 | 0.3 |
|  | 5 | 28.4 | 19.2 | 11.5 | 19.4 |
|  | 6 | 28.4 | 32 | 33.8 | 35.5 |
|  | 7 | 22.1 | 25 | 32.4 | 32.7 |
|  | 8 | 4.2 | 9.6 | 12.5 | 8.5 |
|  | 9 | 1.1 | 4.4 | 4.4 | 2.4 |
|  | 10 | 2.1 | 0.6 | 1.8 | 0.9 |
|  | 11 |  | 2.3 | 1.2 |  |
|  | 12 | 4.2 | 0.9 | 1.2 | 0.3 |
|  | 13 |  | 1.1 | 0.7 |  |
|  | 14 |  | 0.3 |  |  |
|  | $15+$ |  |  | 0.3 |  |

Table B2.9.2. Age composition data (number of fish aged) from the trawl survey (MacNair et al. 2001).

| Age | Female |  | Male |
| ---: | ---: | ---: | ---: |
| 1 | 1 | 5 |  |
| 2 | 22 | 104 |  |
| 3 | 42 | 178 |  |
| 4 | 55 | 109 |  |
| 5 | 58 | 106 |  |
| 6 | 68 | 92 |  |
| 7 | 41 | 96 |  |
| 8 | 25 | 39 |  |
| 9 | 10 | 22 |  |
| 10 | 10 | 4 |  |
| 11 | 2 | 1 |  |
| 12 | 0 | 1 |  |
| 13 | 1 | 0 |  |

Table B2.10.1. Ratio of discards to retained recreationally caught fish in numbers of fish.

| CPFV |  |  |  | Other |
| :---: | :---: | :---: | :---: | :---: |
| Year | RecFIN | Logbook | Observer | RecFIN |
| 1980 | 3.126219 |  |  | 5.421364 |
| 1981 | 2.739227 |  |  | 8.653592 |
| 1982 | 9.435152 |  |  | 7.000156 |
| 1983 | 9.342324 |  |  | 9.707031 |
| 1984 | 14.5061 |  |  | 13.17026 |
| 1985 | 9.400688 |  |  | 12.85084 |
| 1986 | 7.601768 |  | 2.120172 | 10.37773 |
| 1987 | 4.850019 |  | 2.189922 | 9.266473 |
| 1988 | 7.255059 |  | 2.214286 | 20.41673 |
| 1989 | 6.189065 |  | 3.130631 | 14.245 |
| 1990 |  |  |  |  |
| 1991 |  |  |  |  |
| 1992 |  |  |  |  |
| 1993 | 2.161651 |  |  | 8.828594 |
| 1994 | 2.771594 |  |  | 10.21362 |
| 1995 | 3.017931 | 0.745753 |  | 7.837356 |
| 1996 | 3.362679 | 0.786168 |  | 8.422285 |
| 1997 | 4.416503 | 0.874393 |  | 8.695001 |
| 1998 | 2.694001 | 0.535598 |  | 3.606764 |
| 1999 | 2.400497 | 0.974563 |  | 6.010149 |
| 2000 | 5.674797 | 1.737543 |  | 8.383372 |
| 2001 | 4.241324 | 1.206122 |  | 9.891615 |
| 2002 | 4.154036 | 0.909103 |  | 8.190336 |
| 2003 | 1.582726 | 0.409591 |  | 6.211792 |
| 2004 | 1.349514 | 0.317965 |  | 6.541289 |
| 2005 | 1.316633 | 0.26268 |  | 11.28617 |
| 2006 | 1.820902 | 0.661114 |  | 10.61169 |
| 2007 | 2.842799 | 1.204752 |  | 8.492012 |
| 2008 | 2.073493 |  |  | 10.72115 |
| 2009 | 0.771462 |  |  | 12.00346 |

Table B2.10.2. Discard catch estimates (number of fish observed) used in the model.

| Observer based |  | Logbook based |  |
| :---: | ---: | :---: | ---: |
| Year | Discards | Year | Discards |
| 1986 | 23,352 | 1995 | 6,635 |
| 1987 | 58,745 | 1996 | 8,230 |
| 1988 | 39,965 | 1997 | 4,244 |
| 1989 | 70,189 | 1998 | 2,494 |
|  |  | 1999 | 16,233 |
|  |  | 2000 | 19,035 |
|  |  | 2001 | 10,392 |
|  |  | 2002 | 3,072 |
|  |  | 2003 | 2,488 |
|  |  | 2004 | 1,057 |
|  |  | 2005 | 1,108 |
|  |  | 2006 | 603 |
|  |  | 2007 | 1,106 |

Table B2.10.3. Proportion of the catch, by weight, discarded in the limited entry and open access trawl fisheries with standard errors (from Bellman et al. 2010 and earlier reports). Data is from both north and south of Point Conception.

|  | Limited entry |  | Open access |  |
| ---: | ---: | ---: | :--- | :--- |
| Year | Estimate | SE | Estimate | SE |
| 2002 | 0.3454 | 0.0885 |  |  |
| 2003 | 0.1788 | 0.0271 | 0.1091 | 0.0264 |
| 2004 | 0.0334 | 0.0041 | 0.1008 | 0.0141 |
| 2005 | 0.0266 | 0.0037 | 0.0835 | 0.0142 |
| 2006 | 0.091 | 0.0172 |  |  |
| 2007 | 0.206 | 0.047 | 0.5691 | 0.0723 |
| 2008 | 0.373 | 0.5785 | 0.2602 | 0.2632 |
| 2009 | 0.0749 | 0.1475 | 0.0419 | 0.1268 |

Table B3.3.1. Management quantities for the base case and sensitivities. MSY is the maximum sustainable yield in metric tons. B2011 is the spawning biomass at the start of 2011. B0 is the average spawning biomass in the absence of fishing (virgin spawning biomass). Bmsy is the spawning biomass corresponding to MSY. Fmult is the multiplier on the current fishing mortality (fishing effort) that would produce maximum sustainable yield. $n \operatorname{lnL}$ is the negative loglikelihood of the model fit to the data.

```
Base
Model
```

| MSY | 392 |
| :--- | ---: |
| B2011/BO | 0.14 |
| Bmsy/B0 | 0.07 |
| B2011/Bmsy | 2.19 |
| Fmult | 4.49 |
| nlnL | 3567.46 |

Table B3.4.1. Negative log-likelihood values, current depletion level (B2011/B0), and root mean squared error (RMSE) of the fit to the CPFV index of abundance for the sensitivity analyses.

|  | TOTAL | Survey | Discard | Length comp | Weight comp | Recrui tment | $\begin{aligned} & \text { B2011/B } \\ & 0 \end{aligned}$ | RMSE <br> CPFV |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Base | 3567.46 | -28.89 | 51.20 | 3311.89 | 227.53 | 5.73 | 0.14 | 0.22 |
| Aggregated |  |  |  |  |  |  |  |  |
| CPUE | 3584.15 | -13.91 | 48.29 | 3311.23 | 227.84 | 10.68 | 0.04 | 0.39 |
| h75 (not |  |  |  |  |  |  |  |  |
| converge) | 3571.91 | -28.79 | 51.10 | 3311.76 | 227.49 | 10.35 | 0.05 | 0.22 |
| L2=100 | 3550.80 | -25.61 | 48.57 | 3292.73 | 228.97 | 6.14 | 0.21 | 0.26 |
| $\mathrm{M}=0.30 .4$ | 3541.92 | -23.83 | 50.00 | 3283.03 | 226.97 | 5.75 | 0.25 | 0.28 |
| No reweight | 9183.08 | -25.81 | 50.29 | 8735.76 | 415.81 | 7.02 | 0.16 | 0.26 |



Figure B1.6.1. Commercial catch in metric tons by gear type.


Figure B1.6.2. Recreational catch by fishing mode in thousands of fish.


Figure B2.1.1. Index of relative abundance for the CPFV logbook data.


Figure B2.1.2. Index of relative abundance for the detailed trip targeting CPFV logbook data standardized and unstandardized compared to the aggregated data.


Figure B2.2.1. Index of relative abundance for the CaICOFI data. The lower figure differs from the upper figure only in the range of the $y$-axis.


Figure B2.3.1. Gill net survey indices of relative abundance for the gill net survey data.


Figure B2.4.1. Indices of relative abundance for the impingement and related trawl survey data.


Figure B2.5.1. Scaled indices of relative abundance for the sanitation district trawl data.


Figure B2.6.1. Index of relative abundance for the trawl logbook data. (Note that there was very little data in 1984 when the index is estimated to be zero).


Figure B2.8.1. Commercial fishery average length compositions. (note years differ among fisheries and data types. Male and female data not included in combined and do not sum to one together)


Figure B2.8.2. Recreational fishery average length and weight compositions. (note years differ among fisheries and data types)


Figure B2.8.3. Survey average length compositions. (note years differ among surveys) Lengths less than 10 cm are grouped at 10 cm .


Figure B2.8.4. Reweighting of the composition sample size in number of fish.


Figure B2.10.1. Ratio of discards to retained recreational catches in numbers of fish.


Figure B2.10.2. Proportion of the catch in weight discarded in the limited entry and open access trawl fisheries with 95\% confidence intervals (from Bellman et al. 2010) and earlier reports). Data is from both north and south of Point Conception due to the lack of data in the south and the inability to separate some of the data.


Figure B2.10.3. Length composition for halibut trawl reconstructed from Figure 1 of Bellman et al. (2010). Data is from 2003-2010 and includes 341 individuals. Proportion refers to the proportion that length bin comprises of the total discards.


Figure B3.1.1a. Fit of the model to the CPFV index of relative abundance.

Index BT


Figure B3.1.1b. Fit of the model to the trawl logbook index of relative abundance (note that data is not used to estimate parameters)

Index CaICOFI


Figure B3.1.1c. Fit of the model to the CaICOFI index of relative abundance (note that data is not used to estimate parameters)


Figure B3.1.2. Fit to the CPFV discard data.


Figure B3.2.1. Estimated spawning biomass.

Age-0 recruits (1,000s)


Figure B3.2.2. Recruitment estimates.


Figure B3.2.3. Relationship between recruitment and spawning stock size.

Female ending year selectivity for BT


Figure B3.2.4a. Bottom trawl (BT) female selectivity. (Also used for the single-rigged trawl and trawl with a footrope less than 8 inches)


Figure B3.2.4b. Bottom trawl male selectivity. (Also used for the single-rigged trawl and trawl with a footrope less than 8 inches)

Female ending year selectivity for HL


Figure B3.2.4c. Hook-and-line (HL) female selectivity.


Figure B3.2.4d. Hook-and-line male selectivity.

Female ending year selectivity for GN


Figure B3.2.4e. Gill net (GN) female selectivity.


Figure B3.2.4f. Gill net male selectivity.

Female ending year selectivity for RecPa


Figure B3.2.4g. CPFV (RecParty) female selectivity. (Also used for the other recreational fishery)

Male ending year selectivity for RecPart!


Figure B3.2.4h. CPFV (RecParty) male selectivity. (Also used for the other recreational fishery)

Spawning depletion with $\sim 95 \%$ asymptot



Figure B3.4.1. Model fit to the aggregated CPFV index that is not filtered for targeting.

## Appendix B.1: Fit to the length composition data.

Pearson residuals, sexes combined, reté


Pearson residuals, female, retained, BT


Pearson residuals, male, retained, BT (n


Pearson residuals, sexes combined, reté


Pearson residuals, female, retained, HL ।


Pearson residuals, male, retained, HL (m


Pearson residuals, sexes combined, whc


Pearson residuals, sexes combined, reté


Pearson residuals, female, whole catch,


Pearson residuals, female, retained, GN


Pearson residuals, male, whole catch, G


Pearson residuals, male, retained, GN ( n


Pearson residuals, sexes combined, retá


Pearson residuals, female, retained, SR1


Pearson residuals, male, retained, SRT (


Pearson residuals, sexes combined, retá


Pearson residuals, female, retained, TF 1


Pearson residuals, male, retained, TF (m


Pearson residuals, sexes combined, whc


Pearson residuals, sexes combined, reta


Pearson residuals, sexes combined, retá


Pearson residuals, sexes combined, retá


Pearson residuals, sexes combined, reté

length comps, sexes combined, $n$

length comps, sexes combined, retained

length comps, female, whole catch, ag!

length comps, female, retained, aggrega

length comps, male, whole catch, aggrı

length comps, male, retained, aggregate


